

# Using V-Dem the Right Way: Monte Carlo Techniques for Regressing Random Variables

Dan Pemstein

**NDSU** NORTH DAKOTA  
STATE UNIVERSITY



- (Re)familiarize you with Monte Carlo methods for estimating functions of random variables, integrating/marginalizing.
- (Re)familiarize you with how to work with the output of Markov chain Monte Carlo (MCMC) simulations.
- Introduce the V-Dem measurement model.
- Explain how to incorporate measurement uncertainty in V-Dem variables into statistical analyses (regressing random variables).

If we can sample many times from the density,  $f(\theta)$ , of a random variable,  $\theta$ , we can learn anything we want to know about any computable function of that variable.

- $E(\theta) = \int \theta f(\theta) d\theta$ .
- What if this integral is tricky to compute, but we can sample from  $f(\theta)$ ?
- Sample  $\theta^{(t)}$  for  $t = 1, 2, \dots, T$  draws from  $f(\theta)$ .
- $\sum_{t=1}^T \theta^{(t)} / T \rightarrow \int \theta f(\theta) d\theta$  as  $T \rightarrow \infty$ .

$$x_1 \sim \mathcal{N}(\mu_1, \sigma_1) \quad x_2 \sim \mathcal{N}(\mu_2, \sigma_2)$$

- What's the mean of  $y = x_1 + x_2$ . What's the SD?

- $y \sim \mathcal{N}(\mu_1 + \mu_2, \sqrt{\sigma_1^2 + \sigma_2^2})$

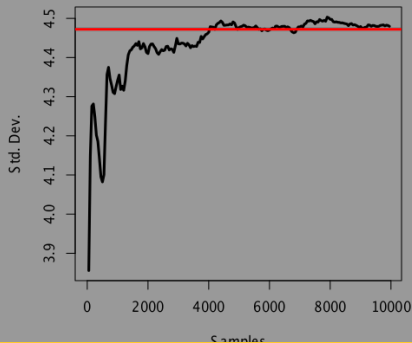
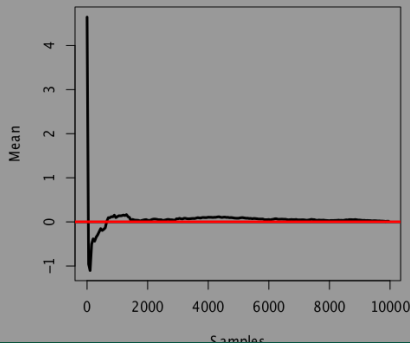
- Simulate:

```
> T <- 10000
> MEAN <- c(-3, 3)
> SD <- c(2, 4)
> x1 <- rnorm(T, MEAN[1], SD[1])
> x2 <- rnorm(T, MEAN[2], SD[2])
> mean(x1 + x2)
[1] 0.01308404
> sd(x1 + x2)
[1] 4.476329
```

# Example: Sums of Random Normal Variables

$$x_1 \sim \mathcal{N}(\mu_1, \sigma_1) \quad x_2 \sim \mathcal{N}(\mu_2, \sigma_2)$$

- What's the mean of  $y = x_1 + x_2$ . What's the SD?
  - $y \sim \mathcal{N}(\mu_1 + \mu_2, \sqrt{\sigma_1^2 + \sigma_2^2})$

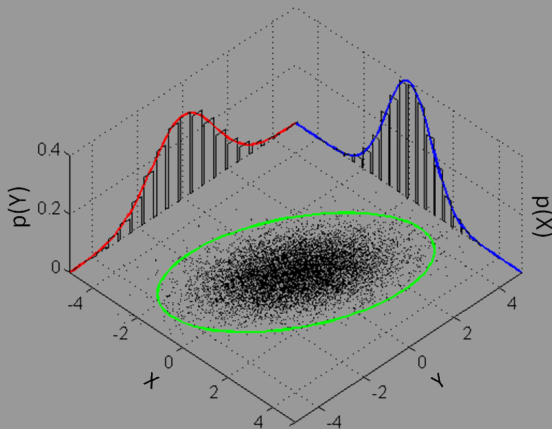


- Say you have a vector of  $n$  random variables  $\boldsymbol{\theta} = \theta_1, \theta_2, \dots, \theta_n$  and data vector  $\mathbf{y}$ .
- The joint posterior density of the random variables is  $f(\boldsymbol{\theta}|\mathbf{y})$ .
- You're interested in the marginal posterior density

$$f(\theta_1|\mathbf{y}) = \int f(\boldsymbol{\theta}|\mathbf{y})d\boldsymbol{\theta}_{-1} = \int f(\theta_1|\boldsymbol{\theta}_{-1}, \mathbf{y})f(\boldsymbol{\theta}_{-1}|\mathbf{y})d\boldsymbol{\theta}_{-1}$$

- If you can sample from  $f(\theta_1|\boldsymbol{\theta}_{-1}, \mathbf{y})$  and  $f(\boldsymbol{\theta}_{-1}|\mathbf{y})$ , then you can simulate from the marginal density  $f(\theta_1|\mathbf{y})$ .
- for each  $t \in 1, 2, \dots, T$  do
  - 1 sample  $\boldsymbol{\theta}_{-1}^{(t)}$  from  $f(\boldsymbol{\theta}_{-1}|\mathbf{y})$
  - 2 sample  $\theta_1^{(t)}$  from  $f(\theta_1|\boldsymbol{\theta}_{-1}^{(t)}, \mathbf{y})$
- $\theta_1^{(t)} \sim f(\theta_1|\mathbf{y})$ .

# The Method of Composition

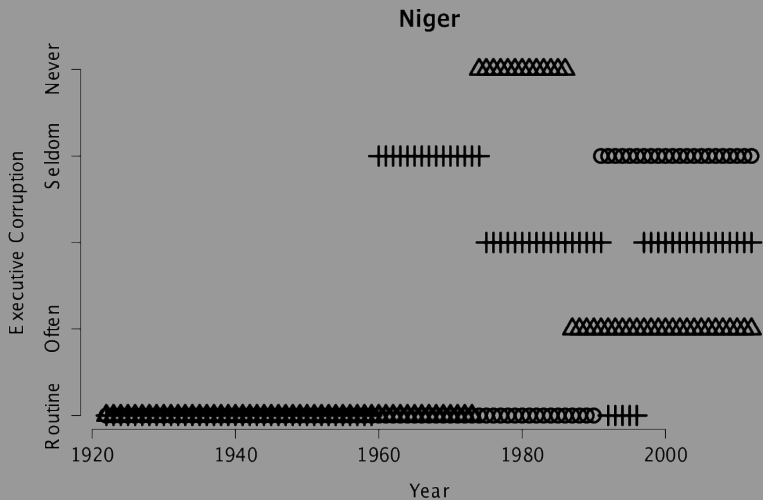


- We want to measure continuous latent traits (matrix of random variables,  $\mathbf{Z}$ ).
- Coders have varying thresholds and reliabilities (coder parameters, vector of random variables,  $\phi$ ).
- Multiple coders per country-year provide observations on an ordinal scale (the data matrix,  $\mathbf{R}$ ).
- We use MCMC methods to simulate latent traits from the marginal posterior density  $f(\mathbf{Z}|\mathbf{R})$ .
- We obtain a sample  $\mathbf{Z}^{(1)}, \mathbf{Z}^{(2)}, \dots, \mathbf{Z}^{(T)}$ :

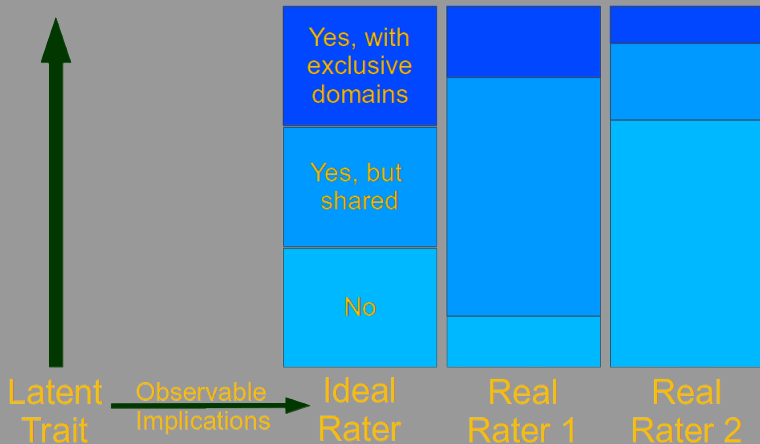
Afghanistan	1900	$z_{11}^{(1)}$	$z_{11}^{(2)}$	$\dots$	$z_{11}^{(T)}$
Afghanistan	1901	$z_{12}^{(1)}$	$z_{12}^{(2)}$	$\dots$	$z_{12}^{(T)}$
$\vdots$	$\vdots$	$\vdots$	$\vdots$	$\vdots$	$\vdots$



- Known only up to a density function.
- Working with point estimates throws out information.
  - This is true for both right and left-hand-side variables.
  - Hard to predict how measurement uncertainty will affect inferences.
    - Cross-correlations in draws.
    - Correlations with other variables may be robust across density.
- Standard “errors in variables” (EIV) model addresses a related, but different issue.
  - Data points in EIV model are country-years in  $\mathbf{R}$ .
  - Our measurement model addresses the EIV problem, while relaxing EIV assumptions about bias (a little bit).
- V-Dem point estimates are best estimates of latent values, but one shouldn’t throw out our uncertainty around those estimates.

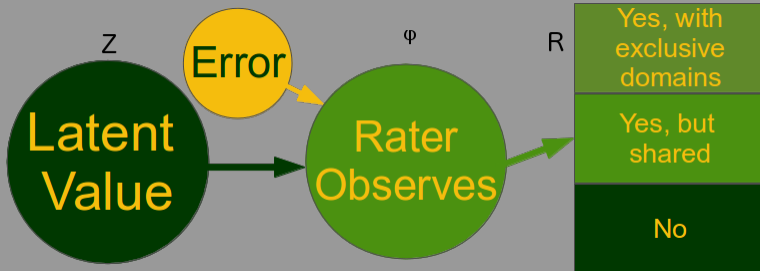


Can the head of government propose legislation in practice?



Multi-rater ordinal probit / ordinal item response theory model

- Rater perceptions equal reality + error
- Raters exhibit arbitrary thresholds
- Given thresholds, raters are correct on average (unbiased)
- Both thresholds and reliability vary across raters



Goal: For an arbitrary regression model, estimate the marginal posterior density of the coefficient vector  $\beta$ , using V-Dem data, taking measurement uncertainty into account.

- Sample from the joint posterior density  $f(\beta, \mathbf{Z} | \mathbf{Y}, \mathbf{R})$ , using the method of composition.
  - $\beta$  is a vector of model coefficients.
  - $\mathbf{Y}$  is a matrix of data measured without uncertainty.
- We can take advantage of the decomposition  $f(\beta, \mathbf{Z} | \mathbf{Y}, \mathbf{R}) = f(\beta | \mathbf{Z}, \mathbf{Y}, \mathbf{R}) f(\mathbf{Z} | \mathbf{R}, \mathbf{Y})$ .
- Assume:
  - $f(\beta | \mathbf{Z}, \mathbf{Y}, \mathbf{R}) = f(\beta | \mathbf{Z}, \mathbf{Y})$ ,
  - $f(\mathbf{Z} | \mathbf{R}, \mathbf{Y}) = f(\mathbf{Z} | \mathbf{R})$ .
- We can now rewrite the decomposition as  $f(\beta, \mathbf{Z} | \mathbf{Y}, \mathbf{R}) = f(\beta | \mathbf{Z}, \mathbf{Y}) f(\mathbf{Z} | \mathbf{R})$ .

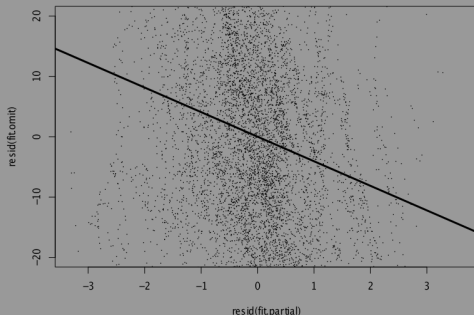
$$f(\beta, \mathbf{Z}|\mathbf{Y}, \mathbf{R}) = f(\beta|\mathbf{Z}, \mathbf{Y})f(\mathbf{Z}|\mathbf{R})$$

- We want the marginal distribution of  $\beta$ ,  $f(\beta|\mathbf{Y})$ .
- Remember:  
$$f(\theta_1|\mathbf{y}) = \int f(\theta|\mathbf{y})d\theta_{-1} = \int f(\theta_1|\theta_{-1}, \mathbf{y})f(\theta_{-1}|\mathbf{y})d\theta_{-1}.$$
- So, given our assumptions:  $f(\beta|\mathbf{Y}) = \int f(\beta|\mathbf{Z}, \mathbf{Y})f(\mathbf{Z}|\mathbf{R})d\mathbf{Z}$ .
- The V-Dem modeling team already simulated  $T=900$  draws where  $\tilde{\mathbf{Z}}^{(t)} \sim f(\mathbf{Z}|\mathbf{R})$  [first MoC step].
- Regression coefficients are distributed  $\mathcal{N}(\mu, \Sigma)$ . To apply the method of composition, for each  $t \in 1, 2, \dots, T$ :
  - 1 Fit your arbitrary regression model to data  $\mathbf{Y}$  and  $\mathbf{Z}^{(t)}$ , yielding partial likelihood estimates  $\hat{\mu}^{(t)}$  and  $\hat{\Sigma}^{(t)}$ .
  - 2 Sample  $\tilde{\beta}^{(t)} \sim \mathcal{N}(\hat{\mu}^{(t)}, \hat{\Sigma}^{(t)})$ .

$$\text{infant mortality}_{cy} = \text{free discussion women}_{cy} + \ln(\text{GDPpc})_{cy}$$

Afghanistan	1900	$\bar{z}_{11}$	$z_{11}^{(1)}$	$z_{11}^{(2)}$	$\cdots$	$z_{11}^{(T)}$
Afghanistan	1901	$\bar{z}_{12}$	$z_{12}^{(1)}$	$z_{12}^{(2)}$	$\cdots$	$z_{12}^{(T)}$
$\vdots$	$\vdots$	$\vdots$	$\vdots$	$\vdots$	$\vdots$	$\vdots$

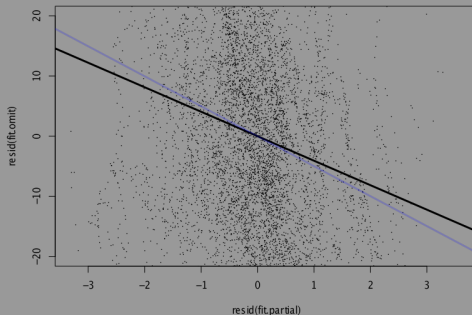
1. Partial regression, point estimate for  $z$  (discussion women).



$$\text{infant mortality}_{cy} = \text{free discussion women}_{cy} + \ln(\text{GDPpc})_{cy}$$

Afghanistan	1900	$\bar{z}_{11}$	$z_{11}^{(1)}$	$z_{11}^{(2)}$	$\dots$	$z_{11}^{(T)}$
Afghanistan	1901	$\bar{z}_{12}$	$z_{12}^{(1)}$	$z_{12}^{(2)}$	$\dots$	$z_{12}^{(T)}$
$\vdots$	$\vdots$	$\vdots$	$\vdots$	$\vdots$	$\vdots$	$\vdots$

2. Fit model using a draw from the marginal density of  $z$ .

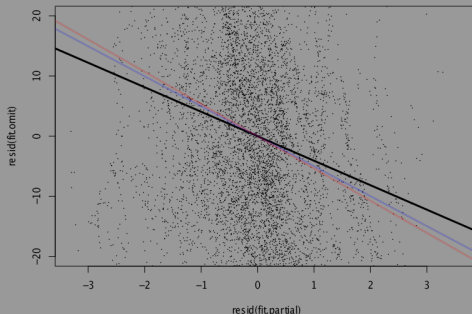




infant mortality<sub>cy</sub> = free discussion women<sub>cy</sub> + ln(GDPpc)<sub>cy</sub>

Afghanistan	1900	$\bar{z}_{11}$	$z_{11}^{(1)}$	$z_{11}^{(2)}$	...	$z_{11}^{(T)}$
Afghanistan	1901	$\bar{z}_{12}$	$z_{12}^{(1)}$	$z_{12}^{(2)}$	...	$z_{12}^{(T)}$
⋮	⋮	⋮	⋮	⋮	⋮	⋮

3. Sample  $\tilde{\beta}^{(t)} \sim \mathcal{N}(\hat{\mu}^{(t)}, \hat{\Sigma}^{(t)})$ .



$$\text{infant mortality}_{cy} = \text{free discussion women}_{cy} + \ln(\text{GDPpc})_{cy}$$

Afghanistan	1900	$\bar{z}_{11}$	$z_{11}^{(1)}$	$z_{11}^{(2)}$	$\dots$	$z_{11}^{(T)}$
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$\vdots$	$\vdots$	$\vdots$	$\vdots$	$\vdots$	$\vdots$	$\vdots$

